

INDEPENDENCE

Durbin-Watson Test

H_0 : not autocorrelations

H_1 : autocorrelations

$$D = \frac{\sum_{i=2}^n (e_i - e_{i-1})^2}{\sum_{i=1}^n e_i^2}$$

THE RANGE IS $0 \leq D \leq 4$

$D = 2$ H_0 IS TRUE

$D < 2$ TOWARDS A POSITIVE AUTOCORRELATION

$D > 2$ TOWARDS A NEGATIVE AUTOCORRELATION

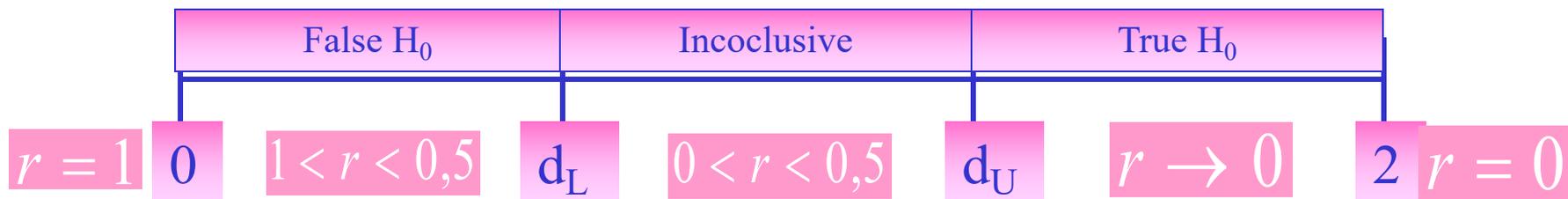
if $D < 2$

H_0 : not autocorrelation
 H_1 : positive autocorrelation

On Durbin-Watson table we read the values of d_L e d_U (where “ n ” is the sample size and “ k ” the number of explanatory variables and $\alpha = 0.05$)

H_1 if $D < d_L$

H_0 if $D > d_U$

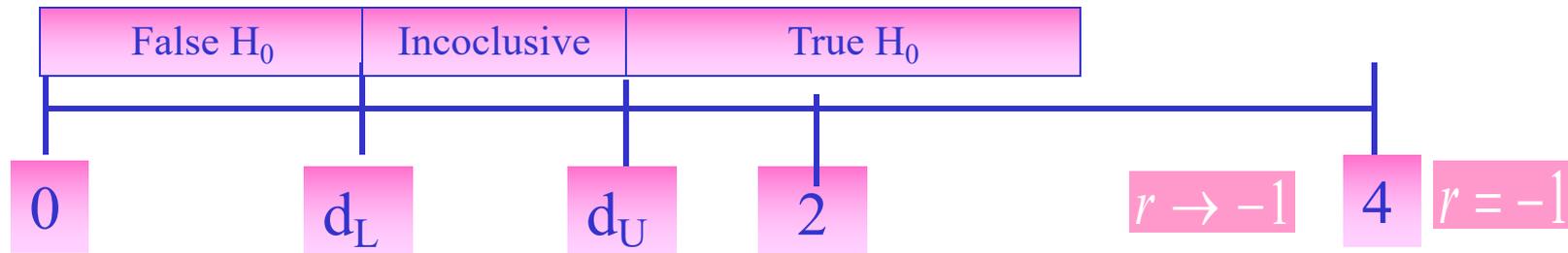


If $D > 2$

H_0 : not autocorrelation
 H_1 : negative autocorrelation

H_1 if $(4-D) < d_L$

H_0 if $(4-D) > d_U$



Durbin-Watson

Y	e _i	e _i ²	(e _i -e _{i-1}) ²	Y*
950	-132.37	17522.03	-	1082.37
1600	-121.37	14730.87	121.00	1721.37
1200	-132.65	17594.91	127.13	1332.65
1500	10.80	116.62	20576.47	1489.20
950	8.21	67.39	6.71	941.79
1700	-58.65	3439.33	4469.59	1758.65
1650	263.04	69189.62	103481.24	1386.96
935	-15.31	234.42	77478.72	950.31
875	-47.62	2267.74	1043.94	922.62
1150	-45.26	2048.54	5.57	1195.26
1400	51.38	2639.82	9339.29	1348.62
1650	104.35	10889.80	2806.35	1545.65
2300	8.85	78.40	9120.25	2291.15
1800	164.89	27190.10	24348.48	1635.11
1400	-28.50	812.01	37399.69	1428.50
1450	-31.75	1007.80	10.56	1481.75
1100	-403.05	162445.92	137863.69	1503.05
1700	182.04	33140.09	342330.31	1517.96
1200	-201.87	40751.82	147390.73	1401.87
1150	18.64	347.42	48624.66	1131.36
1600	-26.59	706.80	2045.30	1626.59
	365.28	133428.89	153558.18	1284.72
	218.80	47875.28	21454.93	981.20
	-442.12	195470.80	436821.86	1242.12
	294.88	86953.74	543169.00	1455.12
		870950.16	2123593.63	

$$D = \frac{\sum_{i=2}^n (e_i - e_{i-1})^2}{\sum_{i=1}^n e_i^2} = \frac{21235993.63}{870950.16} = 2.44$$

Durbin-Watson table

n	$k'=1$		$k'=2$		$k'=3$		$k'=4$		$k'=5$	
	d_L	d_U								
15	1.077	1.361	0.945	1.543	0.814	1.750	0.685	1.977	0.562	2.220
16	1.106	1.371	0.982	1.539	0.857	1.728	0.734	1.935	0.615	2.157
17	1.133	1.381	1.015	1.536	0.897	1.710	0.779	1.900	0.664	2.104
18	1.158	1.392	1.046	1.535	0.933	1.696	0.820	1.872	0.710	2.060
19	1.180	1.401	1.075	1.535	0.967	1.685	0.859	1.848	0.752	2.022
20	1.201	1.411	1.100	1.537	0.998	1.676	0.894	1.828	0.792	1.991
21	1.221	1.420	1.125	1.538	1.026	1.669	0.927	1.812	0.828	1.964
22	1.240	1.429	1.147	1.541	1.053	1.664	0.958	1.797	0.863	1.940
23	1.257	1.437	1.168	1.543	1.078	1.660	0.986	1.786	0.895	1.919
24	1.273	1.446	1.188	1.546	1.101	1.657	1.013	1.775	0.925	1.902
25	1.288	1.454	1.206	1.550	1.123	1.654	1.038	1.767	0.953	1.886
26	1.302	1.461	1.224	1.553	1.143	1.652	1.062	1.759	0.979	1.873
27	1.316	1.468	1.240	1.556	1.162	1.651	1.083	1.753	1.004	1.861
28	1.328	1.476	1.255	1.560	1.181	1.650	1.104	1.747	1.028	1.850
29	1.341	1.483	1.270	1.563	1.198	1.650	1.124	1.743	1.050	1.841
30	1.352	1.489	1.284	1.567	1.214	1.650	1.143	1.739	1.070	1.833

With one regressor and $\alpha=0.05$

Durbin-Watson test

- $D=2.44$
- $n = 25, k = 1$
- $d_L = 1.29$ and $d_U = 1.45$
- $(4-D) = (4-2.44) = 1.56 > d_U = 1.45$, H_0 is true and does not exist negative autocorrelation

True H_0 since

$$(4-D) > d_U$$



Normality

- Histogram
- Box-plot
- Normal Probability Plot (PP PLOT) or QQ Plot
- Statistical Tests (Jarque-Bera, Kolmogorov-Smirnov, Shapiro-Wilk)

Histogram and normality tests

The histogram is a graph which displays the probability distribution of a character (residuals). In this context we want know if the distribution of residuals can be approximated by a normal curve graphically. It is also possible to calculate some statistical tests of hypotheses to be tested, for which:

H_0 : normal

H_1 : not normal

Statistical tests:

Kolmogorov - Smirnov

Jarque - Bera

Shapiro - Wilk

Normality: Kolmogorov-Smirnov (D) test

$$H_0 : F(X) = F_0(X)$$

$$H_1 : F(X) \neq F_0(X)$$

$$D = \sup | \hat{F}_n(X) - F_0(X) |$$

where

$$\hat{F}_n = \frac{1}{n} \sum_i I_{(X_i) \leq X}$$

$F(X)$ is the cumulative distribution function (cdf) of X

In H_0 we refer to the normal cdf.

H_0 is true if the cdfs are similar and, then, the difference is close to zero.

Normality: Shapiro-Wilk (W) test

$$\left\{ \begin{array}{l} H_0: \text{normal distribution} \\ H_1: \text{otherwise} \end{array} \right.$$

$$W = \frac{\left[\sum_{i=1}^n a_i x_i \right]^2}{\sum_{i=1}^n (x_i - \bar{x})^2 a_i}$$

is weight which depends on sample size

$0 < W < 1$, respectively, false and true H_0

The index W is however highly asymmetric.

The threshold values of such statistic are reported in special tables

Normality: Jarque-Bera(JB) test

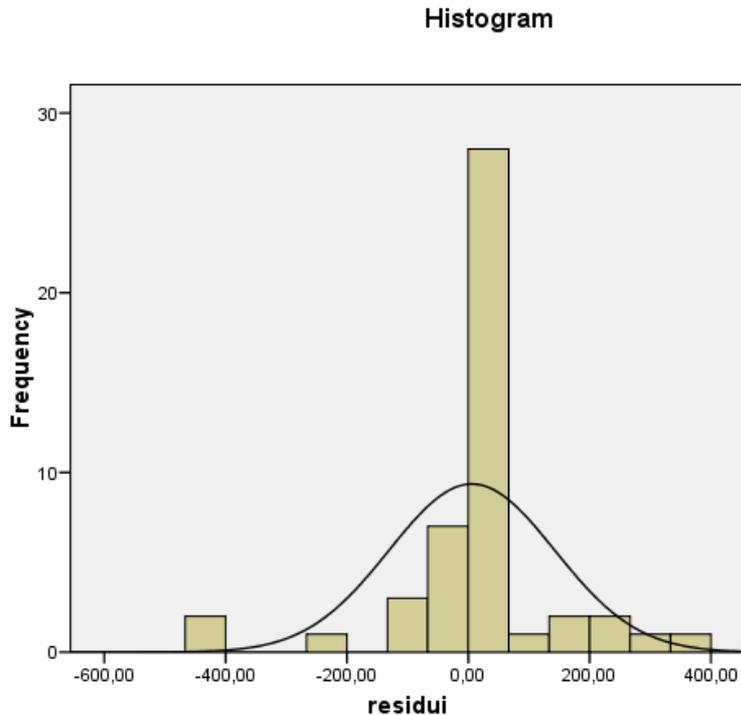
H_0 : normal

H_1 : otherwise

$$JB = n \left[\frac{S^2}{6} + \frac{(K - 3)^2}{24} \right]$$

- **S = Skewness**
- **K = Kurtosis**
- **JB is distributed as χ^2 with 2 d.o.f.**
- **$JB > \chi^2_{2}$ false H_0**

NORMALITY: HISTOGRAM AND TESTS



If Kolmogorov - Smirnov and Shapiro Wilk are lower than 1, we reject H_0 .

Sig. is the p-value, or the probability of rejecting H_0 , when H_0 is true (has to be greater of a given α , to accept H_0).

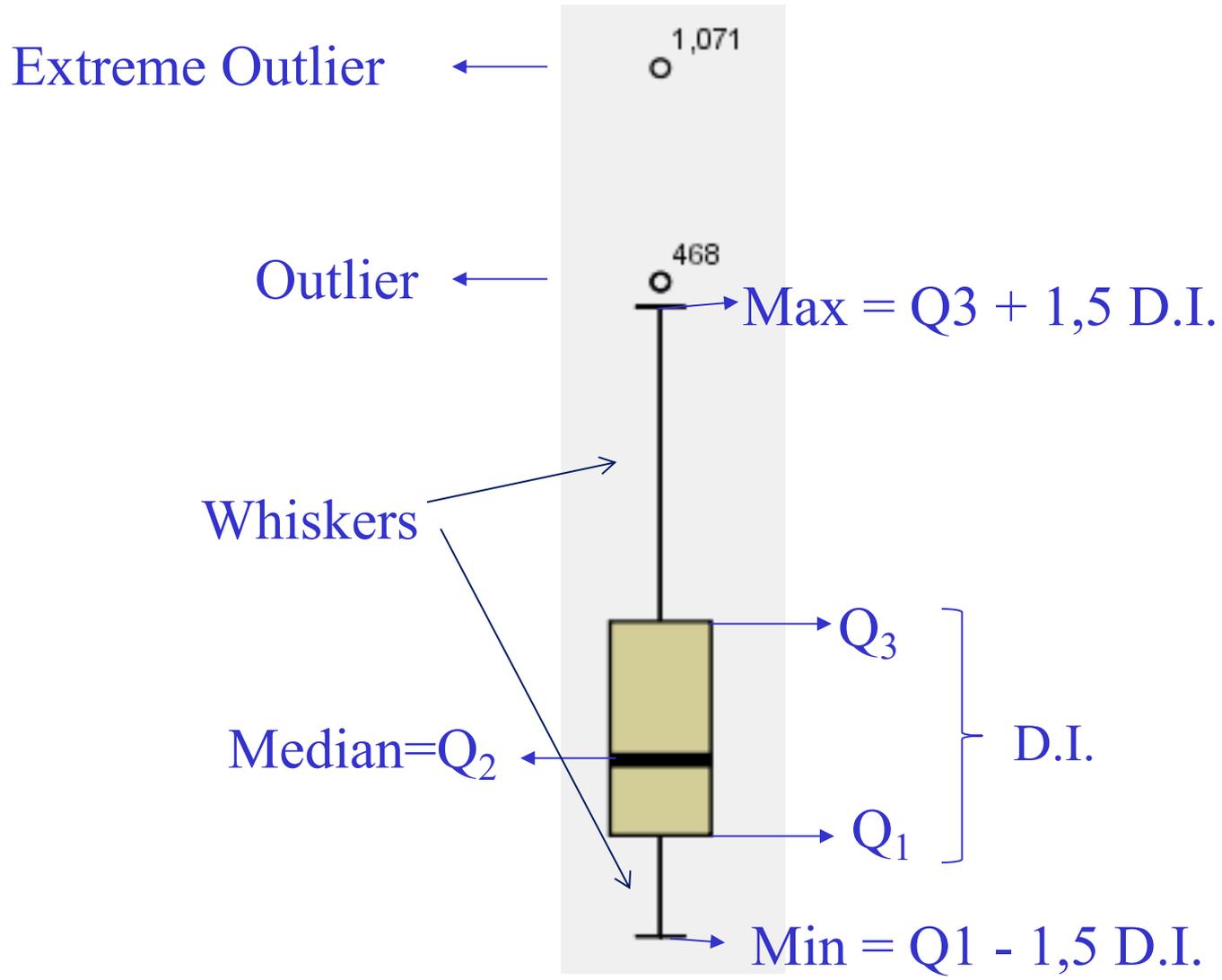
Tests of Normality

	Kolmogorov-Smirnov ^a			Shapiro-Wilk		
	Statistic	df	Sig.	Statistic	df	Sig.
residui	,253	48	,000	,811	48	,000

a. Lilliefors Significance Correction

Not
Normal!

Normality: Box-plot



BOX PLOT:

OUTLIERS

OUTLIERS^o



$$Q_3 + STEP \leq e_i \leq Q_3 + 2STEP$$

or

$$Q_1 - 2STEP \leq e_i \leq Q_1 - STEP$$

Where $STEP = 1.5 * (Q_3 - Q_1) = 1.5 * D.I.$

EXTREME OUTLIERS *



$$e_i < Q_1 - 2STEP$$

or

$$e_i > Q_3 + 2STEP$$

Residui	ni	Ni
-442.12	1	1
-403.05	1	2
-201.87	1	3
-132.65	1	4
-132.37	1	5
-121.37	1	6
-58.65	1	7
-47.62	1	8
-45.26	1	9
-31.75	1	10
-28.50	1	11
-26.59	1	12
-15.31	1	13
8.21	1	14
8.85	1	15
10.80	1	16
18.64	1	17
51.38	1	18
104.35	1	19
164.89	1	20
182.04	1	21
218.80	1	22
263.04	1	23
294.88	1	24
365.28	1	25

EXAMPLE

RANK RESIDUALS!!!

$$\text{position of } Q1 = 25 / 4 = 6.25$$

$$Q1 = -121.37$$

$$\text{position of } Q3 = 25 * 3 / 4 = 18.75$$

$$Q3 = 104.35$$

$$STEP = 1.5 * (Q3 - Q1) = 1.5 * 140.12 = 210.18$$

Outlier

$$Q_3 + STEP \leq e_i \leq Q_3 + 2STEP \rightarrow 104.35 + 210.18 \leq e_i \leq 104.35 + 420.36$$

$$314.53 \leq e_i \leq 524.71$$

or

$$Q_1 - 2STEP \leq e_i \leq Q_1 - STEP \rightarrow -121.37 - 420.36 \leq e_i \leq -121.37 - 210.18$$

$$-541.73 \leq e_i \leq -331.55$$

Limits

$$e_i < Q_1 - 2STEP \rightarrow e_i < -121.37 - 420.36 \rightarrow e_i < -541.73$$

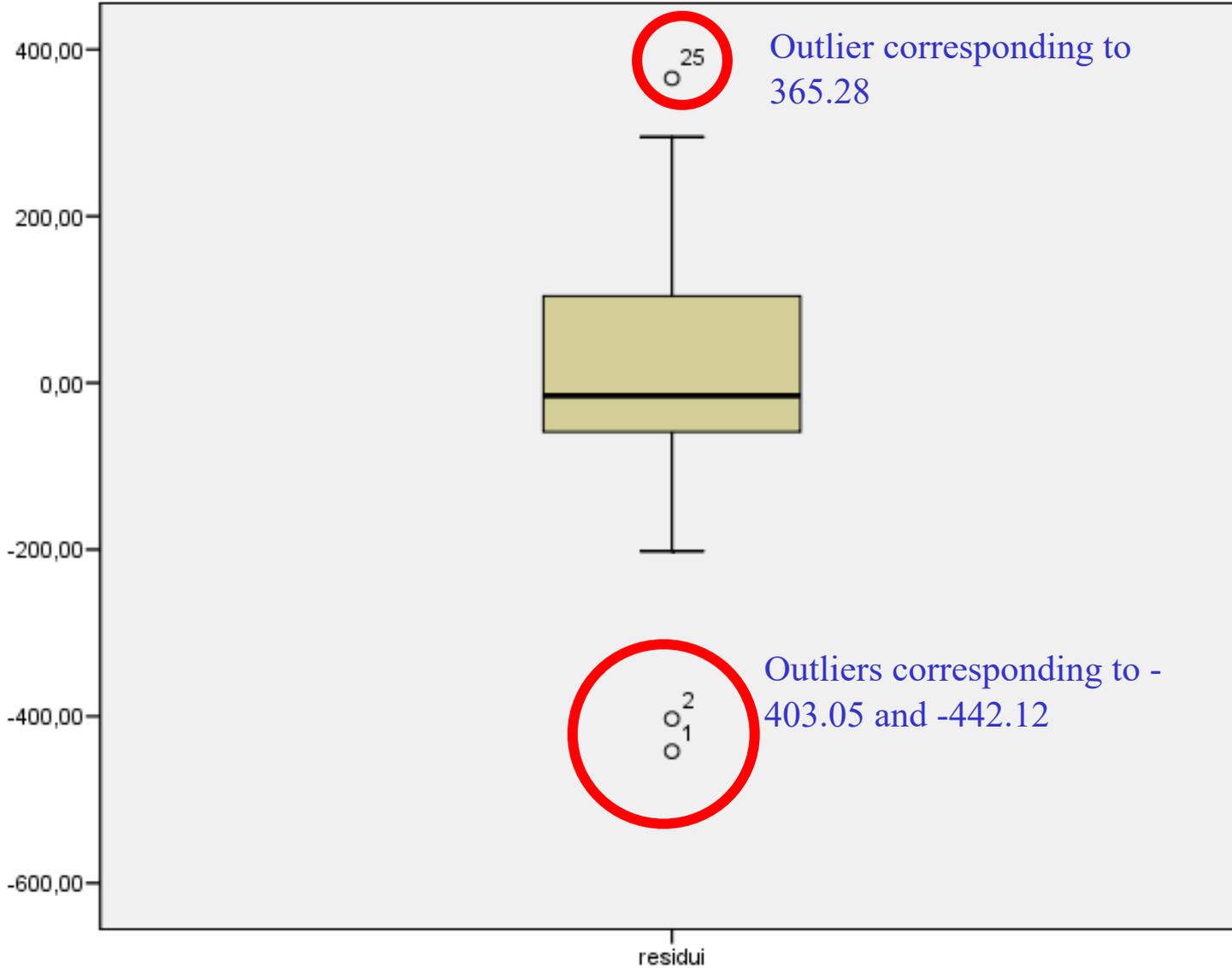
or

or

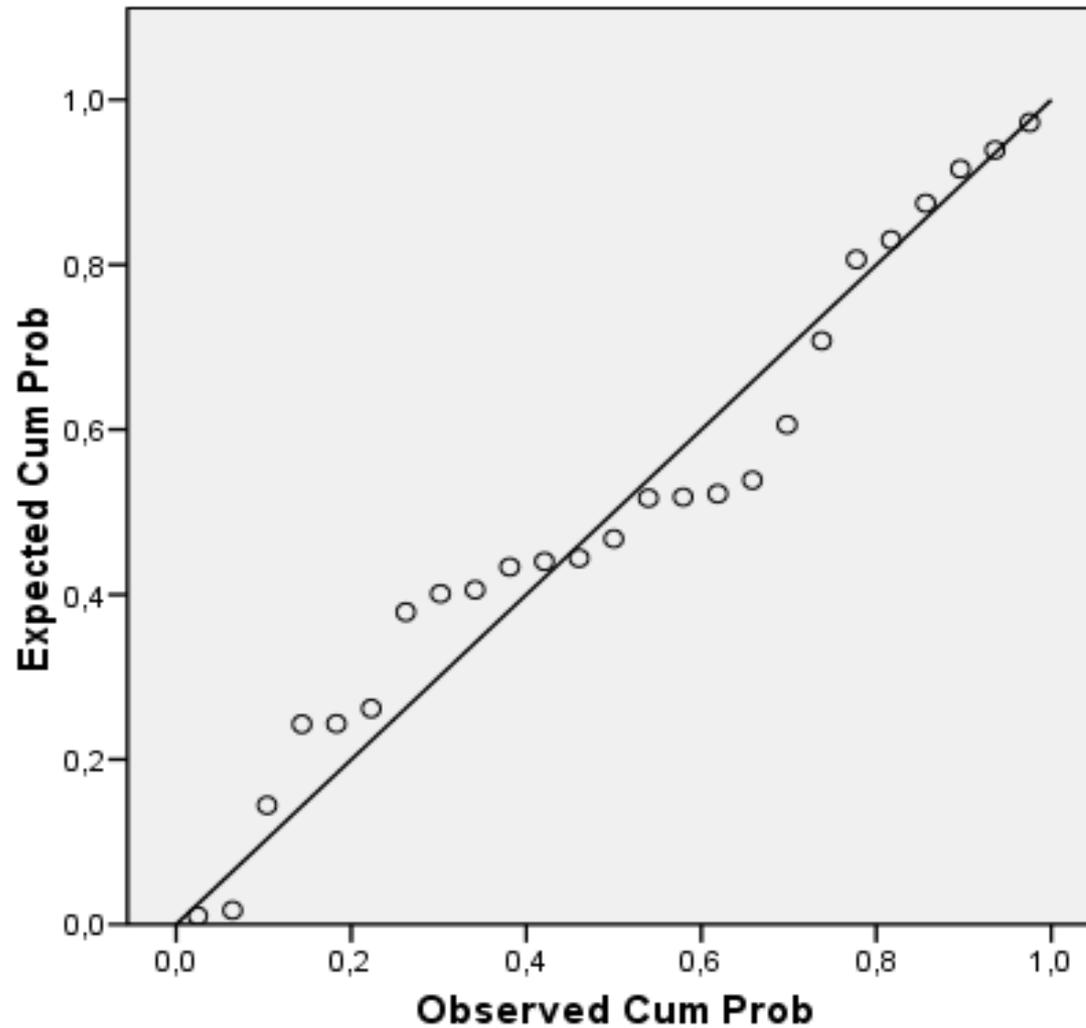
or

$$e_i > Q_3 + 2STEP \rightarrow e_i > 104.35 + 420.36 \rightarrow e_i > 524.71$$

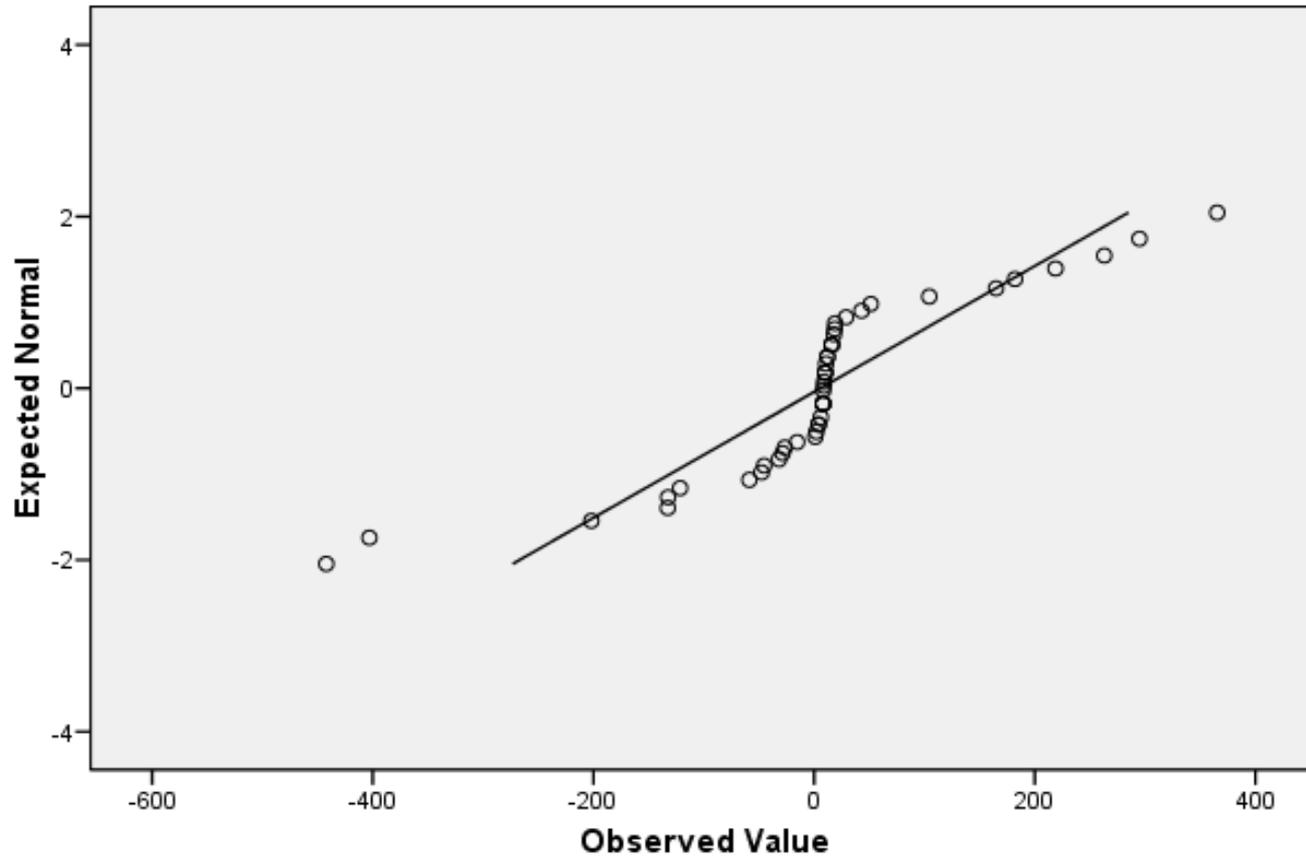
BOX-PLOT



P-P PLOT



Q-Q PLOT

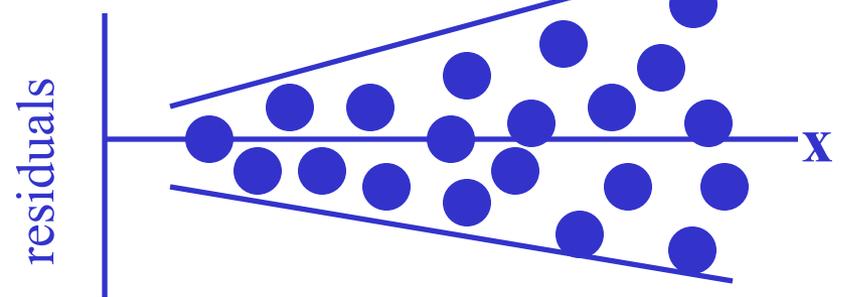
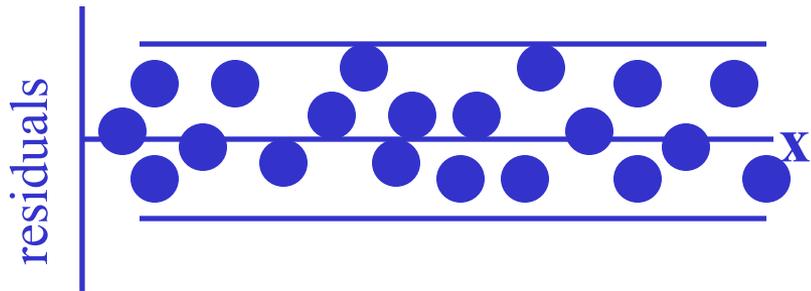
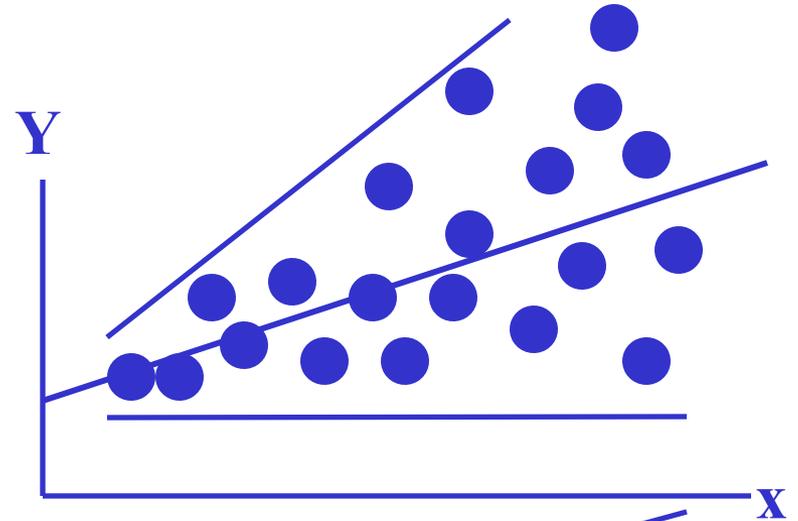
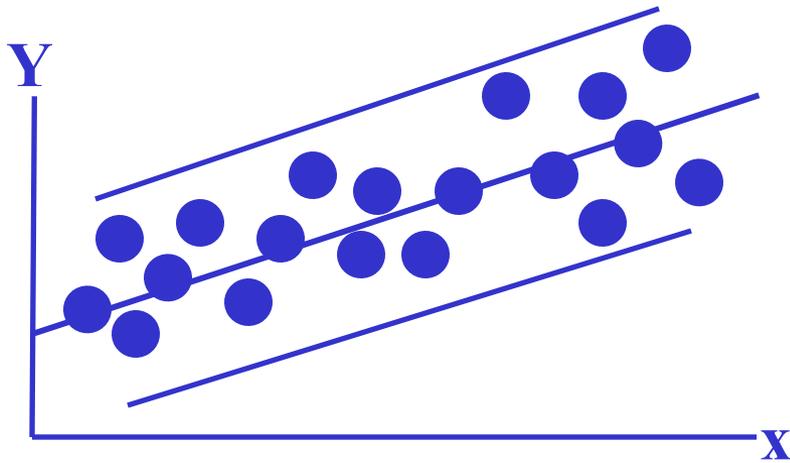


HOMOSKEDASTICITY

FINITE AND CONSTANT VARIANCE OF
RESIDUALS WITH RESPECT TO REGRESSORS

- *Scatter-plot of standardized residuals and regressor (x)*
- *Goldfeld-Quandt test*

Scatter plot: residuals vs x_i



Homoskedasticity

Heteroskedasticity

HOMOSKEDASTICITY:

Goldfeld-Quandt Test

In the case of heteroskedasticity:

- Rank Y with respect to X
- Omit " c " central observations
- Two regressions on the first and the last $(n-c)/2$ observations
- SSE_1 and SSE_2 are the sum of square residuals for the two groups and their ratio is $F_{(n-c-2k)/2}$ and we reject H_0 (Homoskedasticity) if $F_{(n-c-2k)/2} > F_{(n-c-2k)/2, \alpha}$
- In this last case, we can adopt a linear transformation of the dependent variable (Y) or we can apply the method of weighted least squares where the weights are inversely proportional to the error variance